## Moments of Sums

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Let  $X_1, X_2, ..., X_n$  be a sequence of independent random variables. A huge amount of work has been done on estimating the  $L_p$ -norm of the sum of the  $X_s$ :

$$\left\| \sum_{k=1}^{n} X_k \right\|_p = \left\{ E\left( \left| \sum_{k=1}^{n} X_k \right|^p \right) \right\}^{1/p}, \quad p > 0.$$

We first discuss Khintchine's inequality [1], which deals with the Rademacher sequence  $\varepsilon_1, \varepsilon_2, \ldots, \varepsilon_n$ , where

$$P(\varepsilon_k = 1) = P(\varepsilon_k = -1) = 1/2$$
 (symmetric Bernoulli distribution)

for each k. It is known that there exist constants  $A_p$ ,  $B_p$  such that the bounds

$$A_p \left( \sum_{k=1}^n c_k^2 \right)^{1/2} \le \left\| \sum_{k=1}^n c_k \varepsilon_k \right\|_p \le B_p \left( \sum_{k=1}^n c_k^2 \right)^{1/2}$$

hold for arbitrary  $c_1, c_2, \ldots, c_n \in \mathbb{R}$  and  $n \geq 1$ . Szarek [2] and Haagerup [3], building on [4, 5, 6, 7, 8, 9], proved that the best such constants are

$$A_{p} = \begin{cases} ||W||_{p} & \text{if } 0$$

$$B_p = \begin{cases} 1 & \text{if } 0$$

where  $W = 2^{-1/2}(\varepsilon_1 + \varepsilon_2)$ , Z is Normal(0, 1), and  $p_0 = 1.8474163360...$  is the unique solution of the equation

$$\Gamma\left(\frac{p+1}{2}\right) = \frac{\sqrt{\pi}}{2}$$

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in the interval  $0 . In words, if <math>\sum_{k=1}^{n} c_k^2 = 1$ , then  $A_1 = 2^{-1/2}$  and  $B_1 = 1$  encompass the average of  $|\pm c_1 \pm c_2 \pm \cdots \pm c_n|$  taken over all  $2^n$  possible choices of signs. See also [10, 11, 12, 13, 14, 15].

A complex analog of Khintchine's inequality deals with the Steinhaus sequence  $\varepsilon_1, \varepsilon_2, \ldots, \varepsilon_n$ , where  $\varepsilon_k$  is uniformly distributed on the unit circle  $\{z : |z| = 1\}$  for each k. We keep notation identical to before, except that we allow  $c_1, c_2, \ldots, c_n \in \mathbb{C}$ . The best constants  $A_p$ ,  $B_p$  in the inequality

$$A_p \left( \sum_{k=1}^n |c_k|^2 \right)^{1/2} \le \left\| \sum_{k=1}^n c_k \varepsilon_k \right\|_p \le B_p \left( \sum_{k=1}^n |c_k|^2 \right)^{1/2}$$

were conjectured by Haagerup [16] to be

$$A_{p} = \begin{cases} ||W||_{p} & \text{if } 0$$

$$B_p = \begin{cases} 1 & \text{if } 0$$

where  $W = 2^{-1/2}(\varepsilon_1 + \varepsilon_2)$ ,  $Z = 2^{-1/2}(U + iV)$  with U, V independent and Normal(0, 1), and  $p_0 = 0.4756170089...$  is the unique solution of the equation

$$2^{p/2}\Gamma\left(\frac{p+1}{2}\right) = \sqrt{\pi}\left(\Gamma\left(\frac{p+2}{2}\right)\right)^2$$

in the interval  $0 . Here, if <math>\sum_{k=1}^{n} |c_k|^2 = 1$ , then  $A_1 = \sqrt{\pi}/2$  and  $B_1 = 1$  encompass an average taken over all "complex signs" rather than only "real signs" as earlier. Sawa [17] announced that he could verify significant portions of Haagerup's conjecture, but only the case  $p \approx 1$  was published. See also [14, 15, 18, 19]. We mention as well the following result [20, 21] for which p = 1 and p = 1 and p = 1 is the parameter of interest:

$$E\left(\left|\sum_{k=1}^{n} \varepsilon_{k}\right|\right) = \begin{cases} \frac{2}{\pi} \int_{0}^{\infty} \frac{1 - \cos(t)^{n}}{t^{2}} dt & \text{for the real case} \\ \int_{0}^{0} \frac{1 - J_{0}(t)^{n}}{t^{2}} dt & \text{for the complex case} \end{cases}$$

where  $J_0(t)$  is the zeroth Bessel function of the first kind. On the one hand, we have

$$\frac{2}{\pi} \int_{0}^{\infty} \frac{1 - \cos(t)^{n}}{t^{2}} dt = \frac{n!}{2^{n-1} m! (n - m - 1)!} \sim \sqrt{\frac{2n}{\pi}}$$

for the real case, where  $m = \lfloor (n-1)/2 \rfloor$ . On the other hand, the Bessel integral takes on the values 1,  $4/\pi$ , 1.57459723... and 1.79909248... for n = 1, 2, 3 and 4. Keane [22] recently determined that the third value in this list has the following closed-form expression:

$$\frac{1}{8\pi^3}\Gamma\left(\frac{1}{6}\right)^2\Gamma\left(\frac{1}{3}\right)^2 + 48\pi\Gamma\left(\frac{1}{6}\right)^{-2}\Gamma\left(\frac{1}{3}\right)^{-2} = 1.5745972375...$$

but the fourth value still remains open.

We next discuss Rosenthal's inequalities [23]:

$$\left\| \sum_{k=1}^{n} X_{k} \right\|_{p} \le C_{p} \cdot \max \left\{ \left( \sum_{k=1}^{n} \left\| X_{k} \right\|_{p}^{p} \right)^{1/p}, \left\| \sum_{k=1}^{n} X_{k} \right\|_{1} \right\}, \quad p \ge 1$$

for nonnegative random variables and

$$\left\| \sum_{k=1}^{n} X_{k} \right\|_{p} \leq D_{p} \cdot \max \left\{ \left( \sum_{k=1}^{n} \|X_{k}\|_{p}^{p} \right)^{1/p}, \left\| \sum_{k=1}^{n} X_{k} \right\|_{2} \right\}, \quad p \geq 2$$

for symmetric random variables (meaning that the distribution of -X is the same as the distribution of X). A variation of the latter inequality arises if we loosen the restrictive hypothesis "symmetric" to "zero mean"; the constant is then denoted  $E_p$  rather than  $D_p$ . Johnson, Schechtman & Zinn [24] showed that the growth rate of the best constants  $C_p$ ,  $D_p$ ,  $E_p$  is  $p/\ln(p)$  as  $p \to \infty$ ; in contrast, the growth rate for  $B_p$  is only  $\sqrt{p}$ . Subsequent work [25, 26, 27, 28] yielded that

$$C_{p} = \begin{cases} 1 & \text{if } p = 1\\ 2^{1/p} & \text{if } 1$$

where Q is Poisson(1), Z is Normal(0,1), and R, S are independent Poisson(1/2) variables. It is known that  $\|Q\|_m^m = \alpha_m$  and  $\|R - S\|_{2m}^{2m} = \beta_m$  for integer m, where  $\{\alpha_m\}_{m=1}^{\infty} = \{1, 2, 5, 15, 52, 203, \ldots\}$  is the sequence of Bell numbers [29, 30]

$$\alpha_m = \frac{1}{e} \sum_{j=0}^{\infty} \frac{j^m}{j!} = \frac{d^m}{dx^m} \exp\left(\exp(x) - 1\right) \Big|_{x=0}$$

and  $\{\beta_m\}_{m=1}^{\infty} = \{1, 4, 31, 379, \ldots\}$  is the sequence

$$\beta_m = \frac{2}{e} \sum_{k=1}^{\infty} \sum_{j=0}^{\infty} \frac{k^{2m}}{j!(j+k)!2^{2j+k}} = \frac{d^{2m}}{dx^{2m}} \exp\left(\cosh(x) - 1\right) \Big|_{x=0}.$$

Ibragimov & Sharakhmetov [31] conjectured that

$$E_p = \begin{cases} (1 + ||Z||_p^p)^{1/p} & \text{if } 2$$

and proved that this is true when p=2m; further,  $\|Q-1\|_{2m}^{2m}=\gamma_m$  and  $\{\gamma_m\}_{m=1}^{\infty}=\{1,4,41,715,\ldots\}$  is the sequence

$$\gamma_m = \frac{1}{e} \sum_{j=0}^{\infty} \frac{(j-1)^{2m}}{j!} = \frac{d^{2m}}{dx^{2m}} \exp(\exp(x) - x - 1) \Big|_{x=0}.$$

Combinatorial interpretations apply for each of the three sequences:  $\alpha_n$  is the number of partitions of an n-element set into blocks;  $\beta_n$  is the number of partitions of a 2n-element set into blocks, each containing an even number of elements; and  $\gamma_n$  is the number of partitions of a 2n-element set into blocks, each containing more than one element [30].

Define the following Orlicz-type norm:

$$[\Xi]_p = \inf \left\{ \lambda > 0 : \prod_{k=1}^{\infty} \mathbb{E}\left( \left| 1 + \frac{X_k}{\lambda} \right|^p \right) \le e^p \right\}$$

for an arbitrary sequence  $\Xi = \{X_k\}_{k=1}^{\infty}$  of independent random variables, for any p > 0. We mention Latala's inequality [32]:

$$\frac{e-1}{2e^2} \cdot [\Xi]_p \le \left\| \sum_{k=1}^{\infty} X_k \right\|_p \le e \cdot [\Xi]_p$$

which holds either if all the Xs are nonnegative and  $p \geq 1$ , or if all the Xs are symmetric and  $p \geq 2$ . Observe here that the bounds do not depend on p, unlike the earlier inequalities. For the nonnegative case, Hitczenko & Montgomery-Smith [33] improved the left-hand constant  $(e-1)/(2e^2) = 0.116272...$  to  $\xi = 0.154906...$ , where  $\xi$  is the unique positive solution of the equation

$$\sum_{k=0}^{\infty} \frac{(2k+1)^k}{k!} x^k = e.$$

It is not known if this improvement carries over to the symmetric case, nor whether a calculation of best constants is feasible at present.

**0.1.** Addendum. Assuming  $\sum_{k=1}^{n} c_k^2 = 1$ , it is conjectured that the Rademacher sequence satisfies [34, 35, 36, 37, 38]

$$P_n = P\left(\left|\sum_{k=1}^n c_k \varepsilon_k \le 1\right|\right) \ge \frac{1}{2}$$

always. This inequality is provably true if 1/2 is replaced by 3/8 [35] or if all cs are equal [37]. For the latter scenario, we deduce that

$$\lim_{n \to \infty} P_n = \text{erf}\left(1/\sqrt{2}\right) = 0.6826894921...$$

by the normal approximation to the binomial distribution. This constant also appears in [39] with regard to a continued fraction expansion.

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