REVIEW

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Impact of sodium–glucose cotransporter-2 inhibitors in patients with recent versus previous myocardial infarction: a systematic review and meta-analysis

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Abstract

Background Sodium–glucose cotransporter 2 (SGLT2) inhibitors have been included in heart failure (HF) guidelines because of their benefits in reducing mortality and hospitalization rates. However, the timing and benefits of initiating SGLT2 inhibitors in patients after myocardial infarction (MI) remain controversial. Therefore, we aimed to perform a systematic review and meta-analysis comparing SGLT2 inhibitors with placebo in patients with MI.

Methods We performed a systematic review and meta-analysis to determine the impact of SGLT2 inhibitors in patients with recent or previous MI. We systematically searched PubMed, Cochrane, and Embase for RCTs comparing SGLT2 inhibitors versus placebo in patients with MI. The primary outcome was (1) HF hospitalization. In this analysis, we also included the following secondary outcomes: (2) major adverse cardiovascular events (MACE) defined as a composite of cardiovascular (CV) death, MI or stroke; and (3) all-cause mortality. A subgroup analysis was conducted for the primary outcome, comparing patients who had experienced an MI more than 8 weeks prior to study enrolment (previous MI) versus those who had experienced an MI within the preceding 8 weeks (acute MI). Risk ratios (RRs) and 95% confidence intervals (CIs) were pooled with a random effects model.

Results Our meta-analysis included 10 RCTs comprising 22,266 patients, of whom 11,339 (51.2%) had type 2 diabetes. The mean age was 62 years, and the median follow-up was 21 months. According to the pooled analysis, HF hospitalization rates were lower in patients on SGLT2 inhibitors compared with placebo (RR 0.77; 95% CI 0.69, 0.85; p < 0.001)). Differences in MACE were also observed in favor of SGLT2 inhibitors versus placebo (RR 0.88; 95% CI 0.79, 0.97; p = 0.012). There was no statistically significant difference in all-cause mortality between the groups (RR 0.88; 95% CI 0.79, 0.97; p = 0.012).

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CI 0.78, 1.00; p = 0.058). Benefits of SGLT2 inhibitors for the primary outcome were consistent regardless of the timing of last MI, with no treatment by subgroup interaction (p for interaction = 0.56).

Conclusion In this meta-analysis of patients who experienced MI, the administration of SGLT2 inhibitors was associated with lower rates of hospitalization for HF. In addition, the treatment effect of SGLT2 inhibitors was consistent regardless of whether they were started in the recent versus previous MI setting.

Keywords Sodium–glucose cotransporter 2 inhibitors, Myocardial infarction, Cardiovascular risk, Systematic review and meta-analysis

Introduction

Myocardial infarction (MI) remains the main cause of death in the world [1]. The prognosis of MI has improved [2] due to the implementation of early reperfusion, effective pharmacological therapy, evidence-based management of complications, and individualized treatment for specific populations. Despite these advances, there has been a slowdown in improvements in more recent years, with limited new treatment options and a persistent high residual risk of cardiovascular (CV) events after MI [3].

Sodium–glucose cotransporter-2 (SGLT2) inhibitors have emerged as new and widely studied drugs in cardiology because of their positive effects in a wide spectrum of CV and metabolic parameters Recently, SGLT2 inhibitors have been recommended in heart failure (HF) guidelines [4] to reduce cardiovascular mortality and hospitalization caused by HF exacerbations. Additionally, they have proven benefit in patients with chronic kidney disease and type 2 diabetes (T2DM), with or at risk of atherosclerotic CV disease, based on several randomized controlled trials (RCTs) and meta-analyses [5–7].

Previous meta-analyses [8, 9] of patients with acute MI demonstrated a reduction in HF hospitalization with SGLT2 inhibitors but no statistically significant decrease in the other CV outcomes. However, large RCTs testing SGLT2 inhibitors have included subgroup analyses of patients with a history of MI, which warrants further exploration. Moreover, the optimal timing for initiating SGLT2 inhibitors in MI patients remains controversial. Therefore, we aimed to conduct a systematic review and meta-analysis to reassess the efficacy of SGLT2 inhibitors in patients with MI and compare the outcomes between patients with recent versus previous MI.

Methods

This systematic review and meta-analysis was performed and reported according to the Preferred Reporting Items for Systematic Reviews and Meta-Analysis (PRISMA) Statement guidelines and the Cochrane Collaboration Handbook for Systematic Reviews of Interventions guidelines [10, 11] The prospective meta-analysis protocol was uploaded to the International Prospective Register of Systematic Reviews (PROSPERO; CRD42024566070).

Eligibility criteria

There were no restrictions in publication date, status, or language. The inclusion criteria were as follows: (1) RCTs comprising patients with MI, either with acute MI diagnosis or a history of previous MI; (2) comparing SGLT2 inhibitors with placebo; and (3) reporting any prespecified efficacy and safety outcomes. Given that SGLT2 inhibitors seem to provide CV benefits regardless of history of MI [7], including both recent and previous MI patients enhances the applicability of our meta-analysis. We excluded studies that did not report any of the outcomes of interest or that had overlapping patient populations. RCTs of SGLT2 inhibitors versus placebo in a general population, including those with or without T2DM, were included only if they specifically reported outcomes for a subgroup of patients with MI.

Search strategy and data extraction

We systematically searched PubMed/Medline, EMBASE, and Cochrane from database inception to June 2024. The study selection process included an initial review of titles and abstracts, followed by a thorough examination of the full texts of potentially suitable studies. The full search strategy is reported in Supplementary Methods 3. Eight authors (C.F.; N.A.; R.P.; A.P.; W.F.; A.C.; J.F.; A.S.), in pairs, independently and following a doubleblinded model, extracted selected studies, reviewed the main reports and supplementary materials and extracted the relevant information from the included trials. Any discrepancies were resolved by consensus among the authors or with deliberation with other review team members (P.S.; E.K.).

Endpoints and subgroup analysis

The primary endpoint of this meta-analysis was (1) hospitalizations for HF. We also included the following secondary endpoints: (2) all-cause mortality; (3) CV death; (4) major adverse cardiovascular events (MACE), defined as the composite of CV death, MI or stroke; (5) MI recurrence; and (6) stroke. Additionally, we conducted a prespecified subgroup analysis for the primary endpoint, focusing on the effects of the following factors: presence of T2DM, timing of MI (recent vs. previous), and type of SGLT2 inhibitor used (empagliflozin or dapagliflozin). A recent MI was defined as patients having experienced an MI less than 8 weeks prior to either hospitalization or study enrollment. In contrast, patients who had experienced an MI at least 8 weeks before their current hospitalization or study inclusion were classified as having a previous MI. We chose this time frame based on previous literature suggesting that changes in left ventricular volume and function following an MI are typically observed after 8 weeks [12]. Accordingly, we consider it clinically relevant to explore the potential effects of SGLT2 inhibitors both before and after post-MI cardiac remodeling may have occurred. Detailed definitions of the endpoints can be found in Supplementary Methods 4. We performed post hoc sensitivity analysis for the primary outcome stratified by left ventricular ejection fraction (LVEF) < 50% and MI presentation on electrocardiogram, that is, ST-elevation myocardial infarction (STEMI). We also reassessed the primary outcome after excluding trials that were sub-analyses or post hoc in nature.

Quality assessment

Eight authors (C.F.; N.A.; R.P.; A.P.; W.F.; A.C.; J.F.; A.S.), in pairs, independently assessed the risk of bias for each trial using the criteria outlined in the *Cochrane Handbook for Systematic Reviews of Interventions* [11] through the Revised Cochrane risk of bias tool for randomized trials (RoB-2) [13]. Disagreements were resolved by consensus or, if necessary, by consulting a third author (E.K.; P.S.). We assessed the risk of bias according to the following domains: random sequence generation, allocation concealment, blinding of participants and personnel, blinding of outcome assessment, incomplete outcome data, selective outcome reporting, and other biases. We graded each trial as having a high, low, or unclear risk of bias for each domain. We also performed funnel plot analysis and the Egger test to assess publication bias [14].

Statistical analysis

Endpoints were analyzed using a risk ratio (RR) with 95% confidence intervals (CIs). We also computed the hazard ratio (HR) with 95% CIs for a time-to-event sensitivity analysis. We assessed heterogeneity via the Cochrane Q statistic and Higgins and Thompson's I² using a restricted maximum likelihood estimator. Heterogeneity was low if $I^2 = 25\%$, moderate if $I^2 = 50\%$, or high if $I^2 = 75\%$. The random effects model was used once we assumed different effect sizes in the selected population. Our prespecified subgroup interaction was performed using the Q test method, following a null hypothesis of no interaction between groups expressed as a p value. The reported p values are two-sided, and we made no adjustments for multiple testing. We performed statistical analyses using R version 4.3.2 (R Core Team, Vienna, Austria) and the R package meta [15].

Trial sequential analysis

We used TSA 0.9.5.10 Beta software for trial sequential analysis (TSA) to confirm our meta-analysis results. The type of boundary value for the hypothesis test was set to a two-sided test with an alpha value of 5%. Once the cumulative studies in the Z curve cross the conventional monitoring boundary or the futility area, the results are consistent and should be considered reliable evidence [16].

Results

Study selection and baseline characteristics

The study selection process is presented in Fig. 1. The initial search identified 1348 studies (PubMed [n=283], Embase [n = 706], and Cochrane [n = 359]). After title and abstract screening and removal of duplicates, 54 studies remained to be fully reviewed according to the inclusion and exclusion criteria. From these, ten double-blinded, multicenter RCTs and their respective reports were included [17-29] enrolling a total of 22,266 patients, of whom 11,669 (52.4%) were randomized to SGLT2 inhibitors. A full description of the eligibility criteria per study can be found in the Supplementary Table 1. The included participants had a mean age of 62 years, were mostly male (67.4%), and 51.2% had T2DM. The follow-up ranged from 2.8 to 50.4 months, with a median of 21 months. Regarding the intervention, 1 study used canagliflozin (100 mg; 300 mg), 4 used dapagliflozin (10 mg), and the remaining five studies used empagliflozin (10 mg; 25 mg). Table 1 and Supplementary Table 2 present other important characteristics from each study.

Pooled analysis of all studies

Patients treated with SGLT2 inhibitors experienced lower rates of hospitalizations for HF compared to those in the placebo group (RR: 0.77; 95% CI 0.69, 0.85; p < 0.001; I²:0%; Fig. 2), a result that was consistent in the time to first event analysis (HR: 0.75; 95% CI 0.67, 0.85; p < 0.01; I^2 :0%; Supplementary Fig. 1). For all-cause mortality, there appeared to be a trend towards reduction with SGLT2 inhibitors, although the difference was not statistically significant and there was moderate statistical heterogeneity (RR: 0.88; 95% CI 0.78, 1.00; p = 0.058; I^2 : 43%; Fig. 3). This result persisted when the analysis was done as a time to first event (HR: 0.87; 95% CI 0.75, 1.01; p = 0.07; I^2 : 53%; Supplementary Fig. 2).

For MACE outcome, we observed significant benefits in favor of the SGLT2 inhibitor group in terms of the HR (Supplementary Fig. 3; HR 0.87; 95% CI 0.79, 0.97; p = 0.01; I^2 :0%) and RR (RR: 0.88; 95% CI 0.79, 0.97; p = 0.012; I^2 :0%; Fig. 4). However, there was no statistically significant difference between the groups in terms of the risk of recurrent MI (RR: 1.04; 95% CI 0.80, 1.36; p = 0.75; I^2 :68%; Supplementary Fig. 4A); (HR: 1.00; 95%



Fig. 1 PRISMA flow diagram. Legend: PRISMA flow diagram of study screening and selection

CI 0.73, 1.37; p = 0.98; I^2 :73%; Supplementary Fig. 4B), CV death (RR: 0.91; 95% CI 0.75, 1.10; p = 0.32; I^2 :44%; Supplementary Fig. 5A); (HR: 0.87; 95% CI 0.71, 1.07; p = 0.19; I^2 :59%; Supplementary Fig. 5B), and stroke risk (RR: 0.88; 95% CI 0.65, 1.17; p = 0.37; I^2 :0%; Supplementary Fig. 6A); (HR: 0.87; 95% CI 0.64, 1.18; p = 0.37; I^2 :0%; Supplementary Fig. 6B).

Prespecified subgroup analysis

The subgroup analysis demonstrated that patients, both with and without T2DM, benefited from the use of SGLT2 inhibitors, with a significantly lower risk of HF hospitalizations. No significant heterogeneity in the effects of SGLT2 inhibitors was observed between these subgroups (Fig. 5; p for interaction = 0.79). Both patients with previous MI, defined as two or more months since the event, and those with recent MI, defined as less than two months since the event, appeared to benefit from SGLT2 inhibitors, with no evidence of heterogeneity in the treatment effect between subgroups, as shown in Fig. 6; p for interaction = 0.56. Additionally, empagliflozin and dapagliflozin demonstrated similar efficacy in reducing HF hospitalizations compared with placebo (Supplementary Fig. 7; p for interaction = 0.22).

| Study and year | Sam- ple size | SGLT2 inhibitor | Black (n) | Age (y) [†] | Male, n (%) | DM II, n (%) | eGFR (mL/min /1.73 m ²) [†] | STEMI, n (%) | Follow- up (months) |
|---------------------------------|---------------------|-----------------------------------|---------------|-------------------------|------------------|------------------|---|-----------------|---------------------------|
| DAPA MI 2023 [17] | 4017 | Dapagliflozin 10 mg | 23(0.6) | 62.9 | 3210 (79.9) | 0 (0) | 83.4 | 2893 (72) | 11.6 |
| Adel et al. 2022 [18] | 93 | Empagliflozin 10 mg | N/A | 56 | 56 (60.2) | 93 (100) | N/A | 50 (53.7) | 6 |
| EMBODY 2020 [19] | 96 | Empagliflozin 10 mg | N/A | 64.4 | 77 (80.2) | 96 (100) | 65.4 | N/A | 5.5 |
| EMMY 2022 [20] | 476 | Empagliflozin 10 mg | N/A | 57 | 392 (82) | 63 (13) | 92.0 | N/A | 6.5 |
| EMPACT-MI 2024 [21] | 6522 | Empagliflozin 10 mg | 92 (1.4) | 63.6 | 4897 (75) | 2081 (31.9) | 77.8 | 4845 (74.3) | 17.9 |
| EMPAREG OUTCOME 2019 [22, 23] | 7020* | Empagliflozin 10 mg or 25 mg | 357 (5.0)* | 63.1* | 3336 (71.2)* | 7020 (100)* | 74.0* | N/A | 37.2 |
| DECLARE TIMI 58 2018 [24] | 3584 | Dapagliflozin 10 mg | 94 (2.6) | 62 | 2739 (76.4) | 3584 (100) | 88.0 | N/A | 50.4 |
| DACAMI 2023 [25] | 100 | Dapagliflozin 10 mg | N/A | 56 | 83 (83) | 0 (0) | 84.0 | 100 (100) | 2.8 |
| DELIVER + DAPA-HF 2024 [26, 27] | 3731 | Dapagliflozin 10 mg | 100 (2.7) | 68.7 | 2825 (75.7) | 1835 (49.2) | 62.9 | N/A | 27.6 |
| CANVAS 2021 [28, 29] | 10,142* | Canagliflozin 100 mg or 300 mg | 336 (3.3)* | 63.2* | 6509 (64.2)* | 10,142 (100)* | 76.5* | 104 (24.7)* | 43.3 |

Table 1 Baseline patient and study characteristics

Binary data is displayed as a number (%). [†]Mean or Median; *Data reported from entire study population, not only myocardial infarction patients.SGLT2: sodium– glucose cotransporter 2; DM II: type II diabetes; eGFR: estimated glomerular filtration rate; n: number; STEMI: ST-elevation myocardial infarction; y: year



Fig. 2 Forest plot for heart failure hospitalization. Legend: patients with MI events treated with SGLT2i had a lower risk of having HF hospitalization than did those treated with placebo. Abbreviations: CI confidence interval; HF: heart failure; MH: Mantel–Haenszel; MI: myocardial infarction; RR: risk ratio; SGLT2i: sodium–glucose-transporter-2 inhibitors

| Study | Events | SGLT2i Total | l Events | Placebo Total | Weight | RR | 95% | СІ | Risk Ratio MH, Random, 95% Cl |
|------------------------------|------------------------|-----------------|-------------|-------------------------|--------|------|--------------|-----|----------------------------------|
| Adel et al. 2022 | 0 | 45 | 0 | 48 | 0.0% | | | | |
| CANVAS | 35 | 215 | 25 | 159 | 6.3% | 1.04 | [0.65: 1.6 | 61 | |
| DACAMI | 0 | 50 | 0 | 50 | 0.0% | | [0.000,0 | -1 | |
| DAPA-MI | 41 | 2019 | 33 | 1998 | 6.7% | 1.23 | [0.78: 1.9 | 41 | |
| DECLARE TIMI 58 | 153 | 1777 | 187 | 1807 | 20.8% | 0.83 | [0.68: 1.0 | 21 | — — |
| DELIVER + DAPA-HF | 292 | 1830 | 332 | 1901 | 28.3% | 0.91 | [0.79; 1.0 | 51 | |
| EMBODY | 0 | 46 | 0 | 50 | 0.0% | | L, | -1 | <u> </u> |
| EMMY | 3 | 237 | 0 | 239 | 0.2% | 7.06 | [0.37; 135.9 | 2] | |
| EMPACT-MI | 169 | 3260 | 178 | 3262 | 20.7% | 0.95 | [0.77; 1.1 | 71 | |
| EMPAREG OUTCOME | 140 | 2190 | 104 | 1083 | 17.0% | 0.67 | [0.52; 0.8 | 5] | |
| Total (95% CI) | 833 | 11669 | 859 | 10597 | 100.0% | 0.88 | [0.78; 1.0 | 0] | • |
| Heterogeneity: $Tau^2 = 0$. | 0097; Chi ² | $^{2} = 10.60,$ | df = 6 (P = | 0.10); I ² = | = 43% | | - | | |
| Test for overall effect: Z | = -1.90 (P | = 0.058) | ` | ,. | | | | 0.3 | 0.5 1 2 |
| | | | | | | | | | Favors SGLT2i Favors Placeb |

Fig. 3 Forest plot for all-cause mortality. Legend: patients with MI events using SGLT2i had no significant change in all-cause mortality endpoint compared to placebo. Abbreviations: CI: confidence interval; MH: Mantel–Haenszel; MI: myocardial infarction; RR: risk ratio; SGLT2i: sodium–glucose-transporter-2 inhibitors

| | 5 | GLT2i | Placebo | | | | | Risk Ratio | | |
|--------------------------------------|--------------------|------------|-------------|------------|--------|------|--------------|--------------|----------------|--|
| Study | Events | Total | Events | Total | Weight | RR | 95% CI | MH, Rando | om, 95% Cl | |
| DAPA-MI | 68 | 2019 | 72 | 1998 | 9.5% | 0.93 | [0.67; 1.29] | | | |
| DECLARE TIMI 58 | 270 | 1777 | 321 | 1807 | 46.0% | 0.86 | [0.74; 0.99] | | | |
| DELIVER + DAPA-HF | 267 | 1830 | 311 | 1901 | 44.5% | 0.89 | [0.77; 1.04] | | - | |
| Total (95% Cl) | 605 | 5626 | 704 | 5706 | 100.0% | 0.88 | [0.79; 0.97] | | | |
| Heterogeneity: I au ⁻ = 0 | $0; Chi^{-} = 0.3$ | 30, df = 2 | 2(P = 0.86) |); 1- = 0% | D | | | | | |
| l est for overall effect: Z | . = -2.52 (P | = 0.012 |) | | | | | 0.8 | 1 1.25 | |
| | | | | | | | F | avors SGLT2i | Favors Placebo | |

Fig. 4 Forest plot for major cardiovascular events. Legend: patients with MI events treated with SGLT2i had a significant decrease in the MACE endpoint compared with those treated with placebo. Abbreviations: CI: confidence interval; MH: Mantel–Haenszel; MI: myocardial infarction; RR: risk ratio; SGLT2i: sodium–glucose-transporter-2 inhibitors

| Study or | SGLT2i | | Placebo | | | | | Risk Ratio | | |
|--------------------------------|-------------------------|-----------|-------------------|-------------|--------|------|-----------------|-----------------------------|--|--|
| Subgroup | Events | Total | Events | Total | Weight | RR | 95% CI | MH, Random, 95% Cl | | |
| DM II | | | | | | | | | | |
| Adel et al. 2022 | 0 | 45 | 0 | 48 | 0.0% | | | | | |
| CANVAS | 17 | 215 | 17 | 159 | 5.3% | 0.74 | [0.39; 1.40] | ė | | |
| DECLARE TIMI 58 | 81 | 1777 | 114 | 1807 | 28.1% | 0.72 | [0.55; 0.95] | | | |
| EMBODY | 0 | 46 | 1 | 50 | 0.2% | 0.36 | [0.02; 8.67] < | | | |
| EMPACT-MI | 50 | 1035 | 61 | 1046 | 16.3% | 0.83 | [0.58; 1.19] | | | |
| EMPAREG OUTCOME | 74 | 2190 | 54 | 1083 | 18.3% | 0.68 | [0.48; 0.96] | | | |
| Total (95% CI) | 222 | 5308 | 247 | 4193 | 68.3% | 0.73 | [0.61; 0.88] | • | | |
| Heterogeneity: $Tau^2 = 0$; | Chi ² = 0.84 | 4, df = 4 | (P = 0.93); | $I^2 = 0\%$ | | | | | | |
| Test for overall effect: Z = | -3.41 (P • | < 0.001) | | | | | | | | |
| | | | | | | | | | | |
| | 0 | 50 | • | 50 | 0.00/ | 1 00 | 10 45 0 001 | | | |
| DACAMI | 2 | 50 | 2 | 50 | 0.6% | 1.00 | [0.15; 6.82] < | _ | | |
| DAPA-MI | 27 | 2019 | 32 | 1998 | 8.4% | 0.83 | [0.50; 1.39] | | | |
| EMPACT-MI | 68 | 2225 | 92 | 2216 | 22.8% | 0.74 | [0.54; 1.00] | | | |
| Total (95% CI) | 97 | 4294 | 126 | 4264 | 31.7% | 0.77 | [0.59; 0.99] | | | |
| Heterogeneity: $Tau^2 = 0;$ | $Chi^2 = 0.23$ | 5, df = 2 | (P = 0.88); | $l^2 = 0\%$ | | | | | | |
| Test for overall effect: $Z =$ | = -2.01 (P = | = 0.044) | | | | | | | | |
| Total (95% CI) | 319 | 9602 | 373 | 8457 | 100.0% | 0.74 | [0.64: 0.86] | • | | |
| Heterogeneity: $Tau^2 = 0$ | $Chi^2 = 1.1$ | 5 df = 7 | (P = 0.99) | $l^2 = 0\%$ | | • | [010 1, 0100] [| | | |
| Test for overall effect: Z = | -3.95 (P < | < 0.001) | (. <u>5.00</u>), | . 070 | | | 0.3 | 3 0.5 1 2 | | |
| Test for subgroup differer | nces: Chi ² | = 0.07, c | lf = 1 (P = 0 | 0.79) | | | | Favors SGLT2i Favors Placeb | | |

Fig. 5 Forest plot for subgroup analysis in patients with DM II versus non-DM II. Legend: patients in the T2DM subgroup and those in the non-T2DM subgroup had no difference in the primary endpoint. Abbreviations: CI: confidence interval; DM II: diabetes mellitus type II; MH: Mantel–Haenszel; MI: myocardial infarction; RR: risk ratio; SGLT2i: sodium–glucose-transporter-2 inhibitors

Sensitivity analysis and trial sequential analysis

We performed a sensitivity analysis using the leave-oneout method for the outcomes of HF hospitalizations, all-cause mortality, CV death, and MI recurrence. Our primary outcome showed similar results after each trial was sequentially omitted (Supplementary Fig. 8A). On the other hand, after omitting the EMPAREG-OUT-COME trial, our sensitivity analysis revealed a nonsignificant reduction in all-cause mortality (RR: 0.92; 95% CI 0.84 to 1.02; I2=0%; Supplementary Fig. 8B) and CV death (RR: 0.95; 95% CI 0.84 to 1.07; I2:0%; Supplementary Fig. 8C), favoring the SGLT2 inhibitor group. Finally, for the MI recurrence endpoint, after the DECLARE-TIMI 58 trial was omitted, there was a non significant reduction in the risk (RR: 1.20; 95% CI 0.98, 1.47; I2:0%; Supplementary Fig. 8D) of SGLT-2 inhibitor therapy compared with placebo. In our post hoc sensitivity analysis for the primary endpoint in patients with an LVEF < 50% following acute MI, we found no significant difference between SGLT2 inhibitors and placebo (RR: 0.80; 95% CI 0.63, 1.04; $I^2 = 0\%$; Supplementary Fig. 9A). Similarly, for the HF hospitalization outcome in patients with STEMI, SGLT2 inhibitors had no significant effect compared with placebo (RR: 0.85; 95% CI 0.64, 1.12; $I^2 = 0\%$; Supplementary Fig. 9B). Furthermore, our sensitivity analysis after excluding prespecified analyses and post hoc studies was consistent with the overall findings and demonstrated a reduction in HF hospitalization rates (RR: 0.79; 95% CI 0.69, 0.90; I² = 0%; Supplementary Fig. 9C). A TSA was conducted to ensure robust conclusions regarding the primary outcome. The TSA revealed a Z-curve that reached the required information size (RIS) and crossed the significance threshold, indicating a beneficial effect. Moreover, the TSA for subgroups also showed positive results (Supplementary Fig. 10A). A Z curve that exceeded a threshold indicated a benefit in patients with a history of T2DM and MI and those without a history of MI (Supplementary Fig. 10B to D).



Fig. 6 Forest plot for subgroup analysis in patients with recent MI versus previous MI. Legend: Patients with recent MI and previous MI subgroups showed no difference in the primary endpoint. Abbreviations: CI: confidence interval; MH: Mantel–Haenszel; MI: myocardial infarction; rr: risk ratio; SGLT2i: sodium–glucose-transporter-2 inhibitor

This suggests consistent benefits across these subgroups. However, the TSA for the subgroup of patients without a history of T2DM did not reach conclusive results (Supplementary Fig. 10E), and the Z-curve fell short of the RIS for 17,679 patients. TSA revealed a beneficial effect on MACE, but the analysis did not reach the RIS of 12,683 patients (Supplementary Fig. 10F).

Quality assessment and publication bias

The RoB-2 tool was used for quality assessment. Adel et al.'s study was considered at moderate risk of bias [18], whereas the others remained at low risk, as described in the Supplementary Fig. 11. In the funnel plot analysis, studies presented a symmetrical distribution according to weight and converged toward the pooled effect as the weight increased, as described in the Supplementary Fig. 12. Egger's test also revealed no evidence of publication bias (p = 0.82; Supplementary Fig. 12).

Discussion

In this systematic review and meta-analysis of 10 RCTs enrolling 22,266 participants, we compared SGLT2 inhibitors versus placebo in patients with recent or previous MI. Our main results were as follows: (1) SGLT2 inhibitors reduced hospitalizations for HF; (2) this reduction in hospitalizations for HF with SGLT2 inhibitors was observed regardless of the timing of the MI; (3) HF hospitalization rates did not differ significantly between patients with and without T2DM when SGLT2 inhibitors were used; (4) SGLT2 inhibitors were associated with a lower incidence of MACE; and (5) there was no

significant difference in the incidence of MI, CV death, or all-cause mortality between patients treated with SGLT2 inhibitors versus placebo.

Despite minimal SGLT2 expression in the heart, SGLT2 inhibitors significantly improve cardiac function by enhancing sodium handling and contractility, shifting myocardial energy use toward more efficient substrates, and reducing oxidative stress [30]. Recent RCTs have demonstrated that SGLT2 inhibitors provide CV benefits in various populations, being recognized not only for their glucose-lowering effects but also for their pleiotropic benefits, which include anti-inflammatory and plaque-stabilizing properties. These effects are particularly relevant for patients with complex CV conditions, such as multi-vessel coronary disease [31], a history of MI [32], acute MI patients undergoing PCI [33], and HF [34]. Therefore, SGLT2 inhibitors have acquired more space in the cardiology field, having been suggested as a class I recommendation for HF regardless of the ejection fraction by the European Society of Cardiology guidelines (ESC) [35]. SGLT2 inhibitors may also benefit patients without HF who have experienced MI [26] as they seem to reduce in-hospital arrhythmias [36] and contrast-induced acute kidney injury (CI-AKI) [37] in the post-MI setting. These promising effects have driven further exploration of the role of SGLT2 inhibitors in patients with MI and established coronary artery disease.

The main outcome of our meta-analysis underscores the lower rates of HF hospitalizations with SGLT2 inhibitors in patients following MI. A previous study [38] by Jenca et al. demonstrated that, within one year after an MI, 20 to 30% of patients are diagnosed with late-onset HF. The EMMY trial revealed no significant difference in HF hospitalizations with SGLT2 use in acute MI patients (RR 0.76; 95% CI 0.17, 3.34). However, as noted in the study's limitations, the sample size in this trial was insufficient to provide adequate power for hard clinical endpoints. Our meta-analysis aligns with the results of major RCTs: the EMPACT-MI trial [20] by Butler et al. included 6522 post-MI patients at high CV risk, showing that those treated with empagliflozin experienced fewer HF hospitalizations (2.4 events per 100 patient-years) than those receiving a placebo did (3.6 events per 100 patientyears). Notably, HF hospitalization was a secondary outcome, and the primary outcome, a composite of all-cause mortality or HF hospitalization, was not reduced by empagliflozin. Additionally, SGLT2 inhibitors have been shown to attenuate cardiac remodeling after MI by reducing cardiac fibrosis [39, 40]. Therefore, regarding HF hospitalizations, these results demonstrated that these drugs could significantly impact a patient's prognosis after MI. Moreover, MACE was also lower in the SGLT2 inhibitor group (p = 0.012). The DECLARE TIMI 58 trial [24] compared dapagliflozin with placebo in patients with a history of MI and T2DM, and their results were similar to our analysis (HR 0.84 95% CI 0.72, 0.99). These findings underscore the importance of SGLT2 inhibitor therapy for patients with MI, as it may reduce CV outcomes.

Drugs widely known to reduce mortality in patients with HF, such as beta-blockers [41], angiotensin-converting enzyme (ACE) inhibitors and mineralocorticoid antagonists [42], have demonstrated benefits when started in the acute phase of MI, whereas others, such as sacubitril-valsartan, have not [43], raising concerns about the optimal timing for initiating SGLT2 inhibitors after MI at high risk of developing HF. Patients with a history of previous MI may exhibit distinct baseline characteristics and risk profiles compared with those experiencing recent MI, which can lead to variability in treatment responses. However, including patients with both recent and previous MI can increase the generalizability of the findings, as they mirror real-world clinical scenarios where patients often present with diverse histories of CV events. Although patients who experienced recent MI may be at greater risk for developing CV death and HF events, our analysis demonstrated that SGLT2 inhibitors were also effective in lowering HF hospitalization rates in patients with previous MI (p=0.56). This approach highlights the potential importance of early intervention, suggesting that incorporating SGLT2 inhibitors into the treatment plans of patients with a history of prior MI may contribute to improved long-term CV health. Our TSA analysis indicated that our meta-analysis met the RIS, supporting the robustness of our findings.

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A limitation of previous studies is the lack of data regarding the effect of SGLT2 inhibitors in patients without T2DM. Only 3 out of 10 studies in our meta-analysis included patients without T2DM. The DAPA-AMI trial, which included 4017 patients without T2DM [17] and high CV risk, reported similar hospitalization for HF rates in dapagliflozin and placebo groups (1.3% vs 1.6%; RR 0.83 95% CI 0.5, 1.39). These findings were not consistent with those of a subgroup analysis of patients with (HR 0.91 95% CI 0.63, 1.32) and without diabetes (HR 0.68 95% CI 0.50, 0.93) in the EMPACT-MI trial [21]. Therefore, we performed a subgroup analysis of HF hospitalization by comparing patients with and without T2DM. Our results suggested similar efficacy for both subgroups when treated with SGLT2 inhibitors. However, our TSA results suggested that, in patients without -T2DM, the conventional boundary for RIS was not met, requiring a larger sample size in this subgroup to confirm our findings. This suggests that additional data are needed to confirm the positive trend and draw a definitive conclusion for patients without T2DM.

Additionally, despite the high heterogeneity, our metaanalysis indicated a comparable reduction in all-cause and CV mortality between the SGLT2 inhibitor and placebo groups across trials. Our leave-one-out sensitivity analysis showed that excluding the EMPA-REG OUT-COME trial [22, 23] reduced heterogeneity to 0% for both mortality endpoints. This could be attributed to the trial design, which included patients across a broad CV risk spectrum. Since most post-MI patients were classified as high risk, this may have reduced the observable impact of SGLT2 inhibitors on mortality outcomes. Therefore, further large-scale studies are needed to clarify the effects of SGLT2 inhibitors on CV events and all-cause mortality.

In our meta-analysis, risk of recurrent MI was not different between SGLT2 inhibitors and placebo. The contrasting results observed for this outcome can be partially explained by the post hoc analysis of the CANVAS and CREDENCE programs [29] The CANVAS program tested canagliflozin (100; 300 mg) in patients with previous MI and T2DM and high CV risk. Patients treated with canagliflozin had a substantial increase in MI incidence, and the rate of STEMI was also higher among individuals randomized to canagliflozin than among those randomized to placebo (32% vs 15%; p < 0.001), whereas non-ST-elevation myocardial infarction (NSTEMI) was lower (63% vs. 78%, p < 0.001). The exact mechanism by which canagliflozin demonstrated these results remains unknown but may be related to increase in hematocrit and in blood viscosity caused by SGLT2 inhibitors. Nevertheless, the DECLARE TIMI 58 [24] trial highlighted the reduction in overall rates of MI in patients with previous MI treated with SGLT2 inhibitors. A further subgroup analysis revealed significantly lower rates of type 2 MI (HR 0.64 95% CI 0.42, 0.97), although no corresponding reduction was observed for type I MI. This result demonstrates an important advance in therapies for reducing type 2 MI, considering that few medications have presented positive results for this population.

Our meta-analysis has several significant strengths compared with previous meta-analyses, [8, 9] that explored the effects of SGLT2 inhibitors in acute MI. We addressed several gaps in the literature with our metaanalysis. First, we included five new studies with additional 11,062 patients, enhancing the robustness of our data. Second, our TSA confirmed that the RIS was met, providing sufficient evidence to support the benefits of SGLT2 inhibitors. Third, our subgroup analysis comparing patients with recent versus previous MI revealed that SGLT2 inhibitors showed similar benefits in reducing HF hospitalizations in both groups. This finding helps to resolve a key question in the literature regarding the timing of SGLT2 inhibitor initiation post-MI. Fourth, the MACE outcome rates were lower with SGLT2 inhibitors. Despite the limited number of trials addressing this outcome, these results offer valuable insights into the role of SGLT2 inhibitors in CV care. Finally, we included a time to first event analysis, which was consistent with our major finding.

Our meta-analysis has limitations that warrant consideration. The data from the EMPA-REG OUTCOME and CANVAS trials presented in Table 1 are derived from the original studies [22, 28] rather than the secondary analyses [23, 29] used in the statistical report, which may have led to variability in defining our target population by including patients with atherosclerosis and a history of myocardial infarction (MI), rather than exclusively those with MI. Furthermore, the follow-up duration varied significantly across studies, ranging from 2.8 to 50.4 months, highlighting the need for further RCTs with longer follow-up periods. The incidence of HF hospitalizations may also have been underestimated in two trials [17, 21] due to disruptions caused by the COVID-19 pandemic. Although our TSA of HF hospitalizations demonstrated statistically significant benefits for patients with T2DM, it did not yield significant results for patients without T2DM, leaving an important gap in our findings. Another limitation was the variation of drug dosages investigated across the included studies, as outlined in Table 1, which could introduce heterogeneity in the analysis. Additionally, while MACE definitions vary widely across the literature, all studies in our analysis that evaluated this outcome consistently used the same definition, ensuring a more homogeneous assessment. Moreover, the inclusion of two subgroup analysis [23, 24] from two trials not originally dedicated to the post MI population is another limitation, although we did a sensitivity analysis excluding both and the results

remained consistent with the overall analysis. Finally, the inclusion of a post hoc analysis from the CANVAS and CREDENCE programs [29] was restricted to data from CANVAS patients, as relevant outcomes from the CRE-DENCE program were unavailable.

Conclusion

This systematic review and meta-analysis explored the potential efficacy of SGLT2 inhibitors in patients with MI. This therapy was associated with benefits regarding hospitalization for HF and MACE in patients with both recent and previous MI. These findings suggest that SGLT2 inhibitors might be considered not only in patients with acute MI but also in those with atherosclerosis and a history of previous MI.

Abbreviations

| ACE | Angiotensin-converting enzyme |
|----------|--|
| CI | Confidence interval |
| CI-AKI | Contrast-induced acute kidney injury |
| CV | Cardiovascular |
| ESC | European Society of Cardiology |
| HF | Heart failure |
| HR | Hazard ratio |
| LVEF | Left ventricular ejection fraction |
| MACE | Major adverse cardiovascular events |
| MI | Myocardial infarction |
| NSTEMI | Non-ST-elevation myocardial infarction |
| PRISMA | Preferred Reporting Items for Systematic Reviews and |
| | Meta-Analysis |
| PCI | Percutaneous Coronary Intervention |
| PROSPERO | Prospective Register of Systematic Reviews |
| RCT | Randomized controlled trial |
| RIS | Required information size |
| RoB-2 | Revised Cochrane risk of bias tool for randomized trials |
| RR | Risk ratio |
| SGLT2i | Sodium–glucose cotransporter-2 inhibitor |
| SGLT2 | Sodium–glucose cotransporter-2 |
| STEMI | ST-elevation myocardial infarction |
| TSA | Trial sequential analysis |
| T2DM | Type 2 diabetes mellitus |

Supplementary Information

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Additional file 1.

Author contribution

This manuscript reflects the substantial contributions and responsibilities of each listed author. All authors have made significant contributions to the conception and design of the study, as well as to the analysis and interpretation of the data. Specifically, P.S. led the drafting of the manuscript and guided the team through each phase of the project. E.K. played a major role in the statistical analysis, utilizing R Studio software for the data analysis. R.H.F. and L.P. were involved in critically revising and modifying the manuscript, and they have provided final approval of the submitted version. Each author has reviewed and approved the final version of this manuscript, including any revisions that reflect their specific contributions. Furthermore, all authors agree to be personally accountable for their own contributions and are committed to ensuring that any questions regarding the accuracy or integrity of the research, including parts in which they were not directly involved, will be appropriately addressed and resolved. We believe this manuscript meets the highest standards of authorship and scholarly integrity, and we look forward to the review process.

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Data availability

No datasets were generated or analysed during the current study.

Declarations

Competing interests

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